Vitamin A for COVID-19: real-time meta analysis of 18 studies (13 treatment studies and 5 sufficiency studies)

@CovidAnalysis, March 2024, Version 25 https://c19early.org/vameta.html

Abstract

Statistically significant lower risk is seen for recovery and viral clearance. 7 studies from 6 independent teams in 4 countries show statistically significant improvements.

Meta analysis using the most serious outcome reported shows 38% [4-60%] lower risk. Results are similar for Randomized Controlled Trials, higher quality studies, and peer-reviewed studies. Results are consistent with early treatment being more effective than late treatment.

5 sufficiency studies analyze outcomes based on serum levels, showing 85% [54-95%] lower risk for patients with higher vitamin A levels.

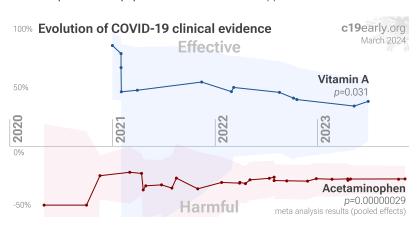
In exclusion sensitivity analysis, statistical significance is lost after excluding 2 of 13 studies in pooled analysis.

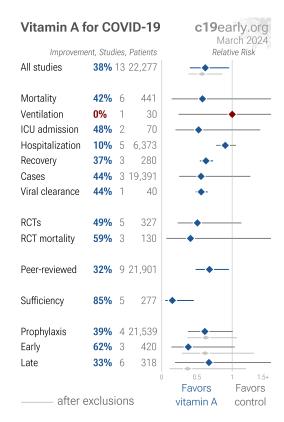
The European Food Safety Authority has found evidence for a causal relationship between the intake of vitamin A and optimal immune system function <code>Galmés</code>, <code>Galmés</code> (B).

No treatment or intervention is 100% effective. All practical, effective, and safe means should be used based on risk/benefit

analysis. Multiple treatments are typically used in combination, and other treatments may be more effective. The quality of non-prescription supplements can vary widely ^{Crawford, Crighton}.

All data to reproduce this paper and sources are in the appendix.





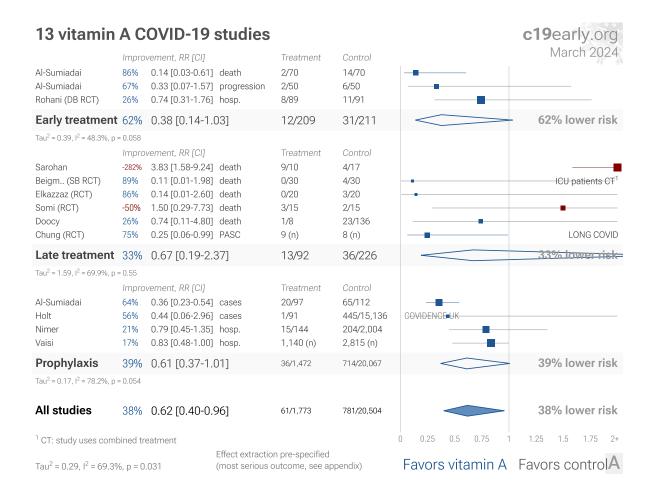
HIGHLIGHTS

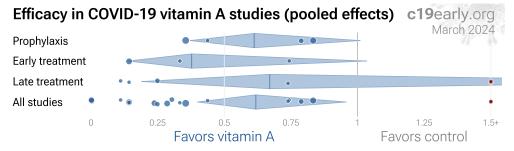
Vitamin A reduces risk for COVID-19 with very high confidence for recovery, high confidence for pooled analysis, low confidence for viral clearance, and very low confidence for ICU admission, hospitalization, progression, and cases.

Vitamin A was the 12th treatment shown effective with \ge 3 clinical studies in January 2021, now known with p = 0.031 from 13 studies.

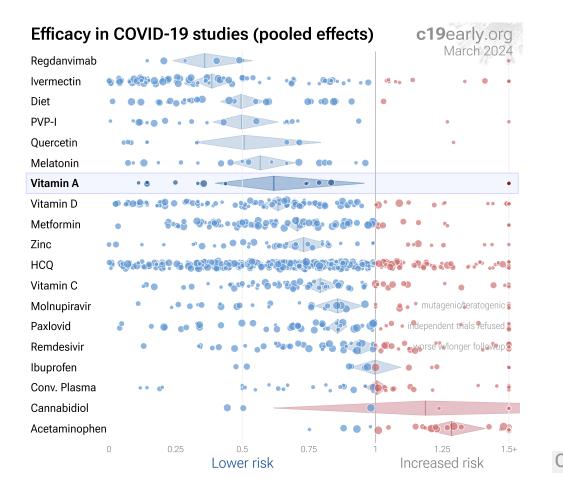
We show traditional outcome specific analyses and combined evidence from all studies, incorporating treatment delay, a primary confounding factor in COVID-19 studies.

Real-time updates and corrections, transparent analysis with all results in the same format, consistent protocol for 66 treatments.





В



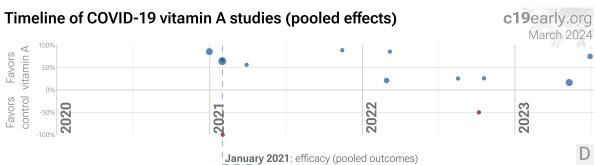


Figure 1. A. Random effects meta-analysis. This plot shows pooled effects, see the specific outcome analyses for individual outcomes, and the heterogeneity section for discussion. Effect extraction is pre-specified, using the most serious outcome reported. For details of effect extraction see the appendix. B. Scatter plot showing the most serious outcome in all studies, and for studies within each stage. Diamonds shows the results of random effects meta-analysis. C. Results within the context of multiple COVID-19 treatments. 0.6% of 6,686 proposed treatments show efficacy c19early.org. D. Timeline of results in vitamin A studies. The marked date indicates the time when efficacy was known with a statistically significant improvement of ≥10% from ≥3 studies for pooled outcomes.

Introduction

Immediate treatment recommended. SARS-CoV-2 infection primarily begins in the upper respiratory tract and may progress to the lower respiratory tract, other tissues, and the nervous and cardiovascular systems, which may lead to cytokine storm, pneumonia, ARDS, neurological issues ^{Scardua-Silva, Yang}, cardiovascular complications ^{Eberhardt}, organ failure, and death. Minimizing replication as early as possible is recommended.

Many treatments are expected to modulate infection. SARS-CoV-2 infection and replication involves the complex interplay of 50+ host and viral proteins and other factors Note A, Malone, Murigneux, Lv, Lui, providing many therapeutic targets for which many existing compounds have known activity. Scientists have predicted that over 6,000 compounds may reduce COVID-19 risk c19early.org (B), either by directly minimizing infection or replication, by supporting immune system function, or by minimizing secondary complications.

Extensive supporting research. Vitamin A has been identified by the European Food Safety Authority (EFSA) as having sufficient evidence for a causal relationship between intake and optimal immune system function ^{EFSA, Galmés, Galmés} (^B). Vitamin A has potent antiviral activity against SARS-CoV-2 in both human cell lines and human organoids of the lower respiratory tract (active metabolite all-trans retinoic acid, ATRA) ^{Tong}, is predicted to bind critical host and viral proteins for SARS-CoV-2 and may compensate for gene expression changes related to SARS-CoV-2 ^{Chakraborty, Huang, Pandya}, may be beneficial for COVID-19 via antiviral, anti-inflammatory, and immunomodulatory effects according to network pharmacology analysis ^{Li}, reduces barrier compromise caused by TNF-α in Calu-3 cells ^{DiGuilio}, inhibits mouse coronavirus replication ^{Franco}, may stimulate innate immunity by activating interferon responses in an IRF3-dependent manner (ATRA) ^{Franco}, may reduce excessive inflammation induced by SARS-CoV-2 ^{Huang}, shows SARS-CoV-2 antiviral activity ^{In Vitro Huang, Moatasim, Morita}, is effective against multiple SARS-CoV-2 variants in Calu-3 cells ^{Morita}, and inhibits the entry and replication of SARS-CoV-2 via binding to ACE2 / 3CLpro / RdRp / helicase / 3′-to-5′ exonuclease Huang.

Analysis. We analyze all significant controlled studies of vitamin A for COVID-19. Search methods, inclusion criteria, effect extraction criteria (more serious outcomes have priority), all individual study data, PRISMA answers, and statistical methods are detailed in Appendix 1. We present random effects meta-analysis results for all studies, studies within each treatment stage, individual outcomes, peer-reviewed studies, Randomized Controlled Trials (RCTs), and higher quality studies.

Treatment timing. Figure 2 shows stages of possible treatment for COVID-19. Prophylaxis refers to regularly taking medication before becoming sick, in order to prevent or minimize infection. Early Treatment refers to treatment immediately or soon after symptoms appear, while Late Treatment refers to more delayed treatment.

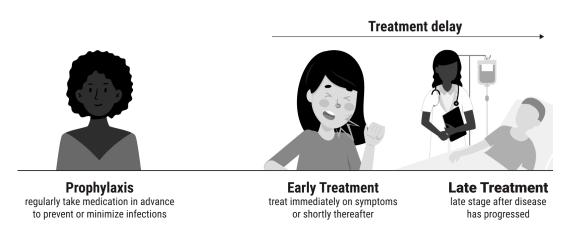


Figure 2. Treatment stages.

Preclinical Research

Vitamin A has potent antiviral activity against SARS-CoV-2 in both human cell lines and human organoids of the lower respiratory tract (active metabolite all-trans retinoic acid, ATRA) ^{Tong}, is predicted to bind critical host and viral proteins for SARS-CoV-2 and may compensate for gene expression changes related to SARS-CoV-2 ^{Chakraborty, Huang, Pandya}, may be beneficial for COVID-19 via antiviral, anti-inflammatory, and immunomodulatory effects according to network pharmacology analysis ^{Li}, reduces barrier compromise caused by TNF-α in Calu-3 cells ^{DiGuilio}, inhibits mouse coronavirus replication ^{Franco}, may stimulate innate immunity by activating interferon responses in an IRF3-dependent manner (ATRA) ^{Franco}, may reduce excessive inflammation induced by SARS-CoV-2 ^{Huang}, shows SARS-CoV-2 antiviral

activity *In Vitro* Huang, Moatasim, Morita, is effective against multiple SARS-CoV-2 variants in Calu-3 cells Morita, and inhibits the entry and replication of SARS-CoV-2 via binding to ACE2 / 3CLpro / RdRp / helicase / 3'-to-5' exonuclease Huang

4 In Silico studies support the efficacy of vitamin A Chakraborty, Huang, Li, Pandya.

5 In Vitro studies support the efficacy of vitamin A DiGuilio, Huang, Moatasim, Morita, Tong.

An In Vivo animal study supports the efficacy of vitamin A Franco.

Preclinical research is an important part of the development of treatments, however results may be very different in clinical trials. Preclinical results are not used in this paper.

Results

Table 1 summarizes the results for all stages combined, for Randomized Controlled Trials, for peer-reviewed studies, after exclusions, and for specific outcomes. Table 2 shows results by treatment stage. Figure 3, 4, 5, 6, 7, 8, 9, 10, 11, 12, and 13 show forest plots for random effects meta-analysis of all studies with pooled effects, mortality results, ventilation, ICU admission, hospitalization, progression, recovery, cases, viral clearance, sufficiency studies, and peer reviewed studies.

	Improvement	Studies	Patients	Authors
All studies	38% [4-60%] *	13	22,277	99
After exclusions	43% [14-62%] **	9	6,739	52
Peer-reviewed studies	32% [5-52%] *	9	21,901	85
Randomized Controlled Trials	49% [-13-77%]	5	327	37
Mortality	42% [-133-86%]	6	441	30
ICU admission	48% [-40-80%]	2	70	11
Hospitalization	10% [-5-23%]	5	6,373	28
Recovery	37% [27-46%] ****	3	280	16
Cases	44% [-26-75%]	3	19,391	42
RCT mortality	59% [-153-93%]	3	130	17
RCT hospitalization	-3% [-27-16%]	3	270	19

Table 1. Random effects meta-analysis for all stages combined, for Randomized Controlled Trials, for peer-reviewed studies, after exclusions, and for specific outcomes. Results show the percentage improvement with treatment and the 95% confidence interval. * p<0.05 *** p<0.01 **** p<0.0001.

	Early treatment	Late treatment	Prophylaxis
All studies	62% [-3-86%]	33% [-137-81%]	39% [-1-63%]
After exclusions	38% [-31-71%]	63% [-20-89%]	38% [-7-63%]
Peer-reviewed studies	64% [-80-93%]	51% [-38-82%]	18% [3-30%] *
Randomized Controlled Trials	26% [-76-69%]	63% [-20-89%]	
Mortality	86% [39-97%] **	9% [-235-75%]	
ICU admission		48% [-40-80%]	
Hospitalization	26% [-76-69%]	3% [-53-39%]	17% [2-30%] *
Recovery	32% [-133-80%]	37% [27-46%] ****	
Cases			44% [-26-75%]
RCT mortality		59% [-153-93%]	
RCT hospitalization	26% [-76-69%]	3% [-53-39%]	

Table 2. Random effects meta-analysis results by treatment stage. Results show the percentage improvement with treatment, the 95% confidence interval, and the number of studies for the stage. *p<0.05 ***p<0.01 **** p<0.0001.

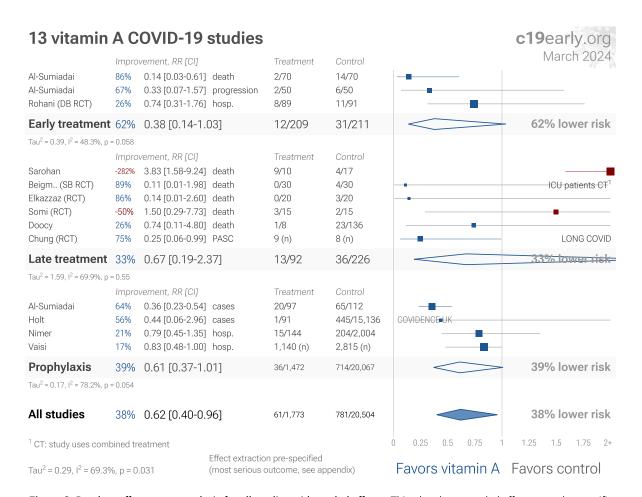


Figure 3. Random effects meta-analysis for all studies with pooled effects. This plot shows pooled effects, see the specific outcome analyses for individual outcomes, and the heterogeneity section for discussion. Effect extraction is pre-specified, using the most serious outcome reported. For details of effect extraction see the appendix.

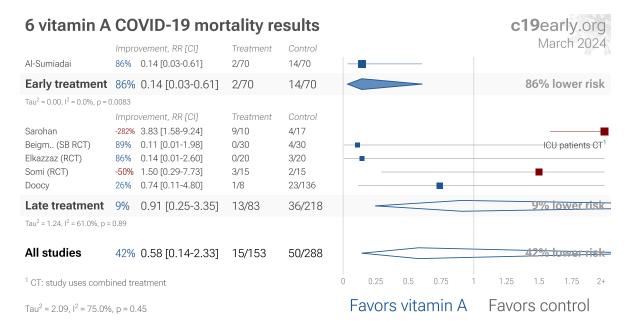


Figure 4. Random effects meta-analysis for mortality results.

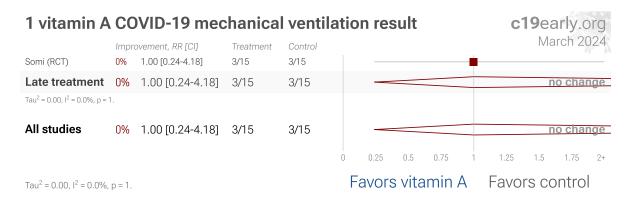


Figure 5. Random effects meta-analysis for ventilation.

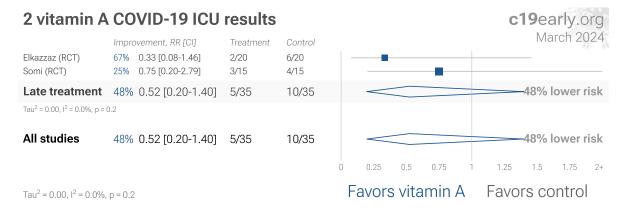


Figure 6. Random effects meta-analysis for ICU admission.

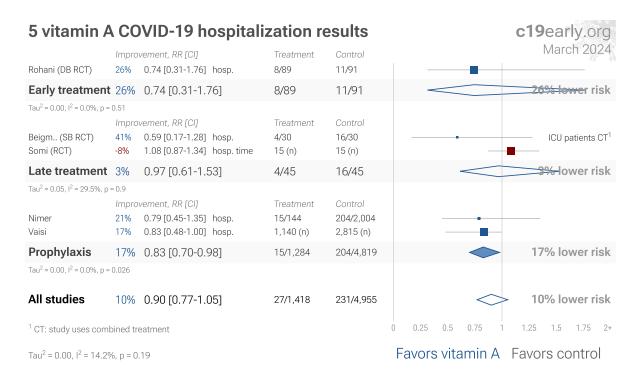


Figure 7. Random effects meta-analysis for hospitalization.

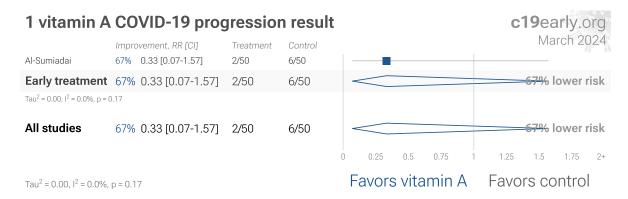


Figure 8. Random effects meta-analysis for progression.

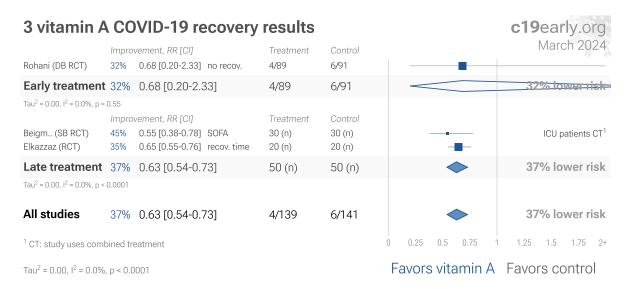


Figure 9. Random effects meta-analysis for recovery.



Figure 10. Random effects meta-analysis for cases.



Figure 11. Random effects meta-analysis for viral clearance.

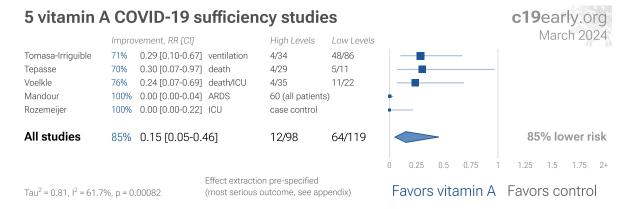


Figure 12. Random effects meta-analysis for sufficiency studies. Effect extraction is pre-specified, using the most serious outcome reported, see the appendix for details.

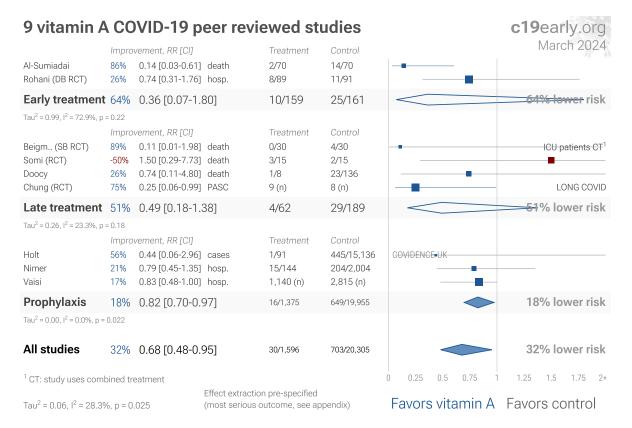


Figure 13. Random effects meta-analysis for peer reviewed studies. Effect extraction is pre-specified, using the most serious outcome reported, see the appendix for details. *Zeraatkar et al.* analyze 356 COVID-19 trials, finding no significant evidence that preprint results are inconsistent with peer-reviewed studies. They also show extremely long peer-review delays, with a median of 6 months to journal publication. A six month delay was equivalent to around 1.5 million deaths during the first two years of the pandemic. Authors recommend using preprint evidence, with appropriate checks for potential falsified data, which provides higher certainty much earlier. *Davidson et al.* also showed no important difference between meta analysis results of preprints and peer-reviewed publications for COVID-19, based on 37 meta analyses including 114 trials.

Randomized Controlled Trials (RCTs)

Figure 14 shows a comparison of results for RCTs and non-RCT studies. The median effect size for RCTs is 75% improvement, compared to 41% for other studies. Figure 15, 16, and 17 show forest plots for random effects meta-analysis of all Randomized Controlled Trials, RCT mortality results, and RCT hospitalization results. RCT results are included in Table 1 and Table 2.

RCTs have many potential biases. Bias in clinical research may be defined as something that tends to make conclusions differ systematically from the truth. RCTs help to make study groups more similar and can provide a higher level of evidence, however they are subject to many biases <code>Jadad</code>, and analysis of double-blind RCTs has identified extreme levels of bias <code>Gøtzsche</code>. For COVID-19, the overhead may delay treatment, dramatically compromising efficacy; they may encourage monotherapy for simplicity at the cost of efficacy which may rely on combined or synergistic effects; the participants that sign up may not reflect real world usage or the population that benefits most in terms of age, comorbidities, severity of illness, or other factors; standard of care may be compromised and unable to evolve quickly based on emerging research for new diseases; errors may be made in randomization and medication delivery; and investigators may have hidden agendas or vested interests influencing design, operation, analysis, and the potential for fraud. All of these biases have been observed with COVID-19 RCTs. There is no guarantee that a specific RCT provides a higher level of evidence.

Conflicts of interest for COVID-19 RCTs. RCTs are expensive and many RCTs are funded by pharmaceutical companies or interests closely aligned with pharmaceutical companies. For COVID-19, this creates an incentive to show efficacy for patented commercial products, and an incentive to show a lack of efficacy for inexpensive

treatments. The bias is expected to be significant, for example *Als-Nielsen et al.* analyzed 370 RCTs from Cochrane reviews, showing that trials funded by for-profit organizations were 5 times more likely to recommend the experimental drug compared with those funded by nonprofit organizations. For COVID-19, some major philanthropic organizations are largely funded by investments with extreme conflicts of interest for and against specific COVID-19 interventions.

RCTs for novel acute diseases requiring rapid treatment. High quality RCTs for novel acute diseases are more challenging, with increased ethical issues due to the urgency of treatment, increased risk due to enrollment delays, and more difficult design with a rapidly evolving evidence base. For COVID-19, the most common site of initial infection is the upper respiratory tract. Immediate treatment is likely to be most successful and may prevent or slow progression to other parts of the body. For a non-prophylaxis RCT, it makes sense to provide treatment in advance and instruct patients to use it immediately on symptoms, just as some governments have done by providing medication kits in advance. Unfortunately, no RCTs have been done in this way. Every treatment RCT to date involves delayed treatment. Among the 66 treatments we have analyzed, 63% of RCTs involve very late treatment 5+ days after onset. No non-prophylaxis COVID-19 RCTs match the potential real-world use of early treatments (they may more accurately represent results for treatments that require visiting a medical facility, e.g., those requiring intravenous administration).

Non-RCT studies have been shown to be reliable. Evidence shows that non-RCT trials can also provide reliable results. *Concato et al.* found that well-designed observational studies do not systematically overestimate the magnitude of the effects of treatment compared to RCTs. *Anglemyer et al.* summarized reviews comparing RCTs to observational studies and found little evidence for significant differences in effect estimates. *Lee et al.* showed that only 14% of the guidelines of the Infectious Diseases Society of America were based on RCTs. Evaluation of studies relies on an understanding of the study and potential biases. Limitations in an RCT can outweigh the benefits, for example excessive dosages, excessive treatment delays, or Internet survey bias could have a greater effect on results. Ethical issues may also prevent running RCTs for known effective treatments. For more on issues with RCTs see *Deaton*, *Nichol*.

Using all studies identifies efficacy 5.7+ months faster for COVID-19. Currently, 44 of the treatments we analyze show statistically significant efficacy or harm, defined as \geq 10% decreased risk or >0% increased risk from \geq 3 studies. Of the 44 treatments with statistically significant efficacy/harm, 28 have been confirmed in RCTs, with a mean delay of 5.7 months. When considering only low cost treatments, 23 have been confirmed with a delay of 6.9 months. For the 16 unconfirmed treatments, 3 have zero RCTs to date. The point estimates for the remaining 13 are all consistent with the overall results (benefit or harm), with 10 showing >20%. The only treatments showing >10% efficacy for all studies, but <10% for RCTs are sotrovimab and aspirin.

Summary. We need to evaluate each trial on its own merits. RCTs for a given medication and disease may be more reliable, however they may also be less reliable. For off-patent medications, very high conflict of interest trials may be more likely to be RCTs, and more likely to be large trials that dominate meta analyses.

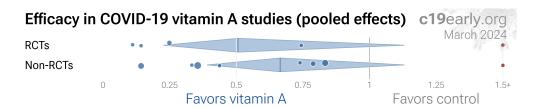


Figure 14. Results for RCTs and non-RCT studies.

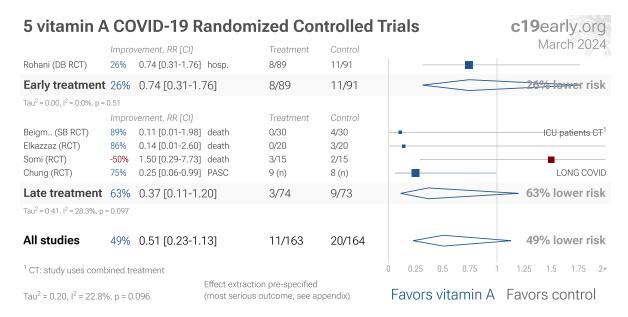


Figure 15. Random effects meta-analysis for all Randomized Controlled Trials. This plot shows pooled effects, see the specific outcome analyses for individual outcomes, and the heterogeneity section for discussion. Effect extraction is prespecified, using the most serious outcome reported. For details of effect extraction see the appendix.

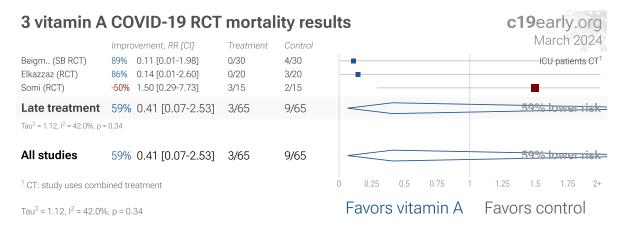


Figure 16. Random effects meta-analysis for RCT mortality results.

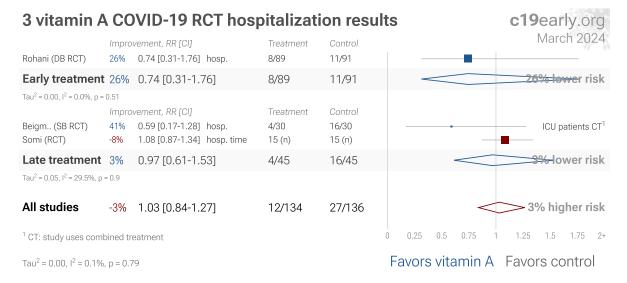


Figure 17. Random effects meta-analysis for RCT hospitalization results.

Exclusions

To avoid bias in the selection of studies, we analyze all non-retracted studies. Here we show the results after excluding studies with major issues likely to alter results, non-standard studies, and studies where very minimal detail is currently available. Our bias evaluation is based on analysis of each study and identifying when there is a significant chance that limitations will substantially change the outcome of the study. We believe this can be more valuable than checklist-based approaches such as Cochrane GRADE, which may underemphasize serious issues not captured in the checklists, overemphasize issues unlikely to alter outcomes in specific cases (for example, lack of blinding for an objective mortality outcome, or certain specifics of randomization with a very large effect size), and can be easily influenced by potential bias.

The studies excluded are as below. Figure 18 shows a forest plot for random effects meta-analysis of all studies after exclusions.

Al-Sumiadai, minimal details of groups provided.

Doocy, unadjusted results with no group details.

Holt, significant unadjusted confounding possible.

Sarohan, unadjusted results with no group details, comments suggest significant group differences and confounding.

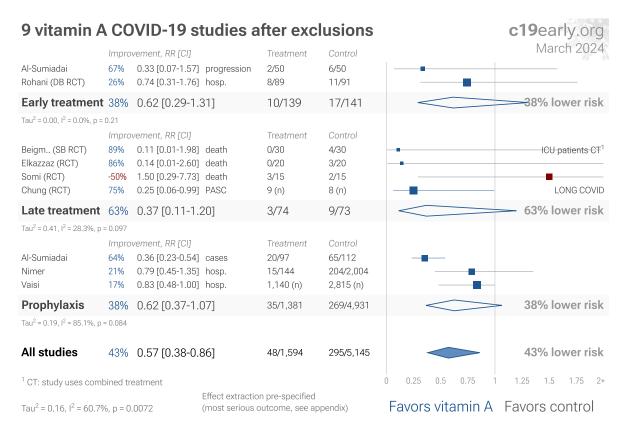


Figure 18. Random effects meta-analysis for all studies after exclusions. This plot shows pooled effects, see the specific outcome analyses for individual outcomes, and the heterogeneity section for discussion. Effect extraction is pre-specified, using the most serious outcome reported. For details of effect extraction see the appendix.

Heterogeneity

Heterogeneity in COVID-19 studies arises from many factors including:

Treatment delay. The time between infection or the onset of symptoms and treatment may critically affect how well a treatment works. For example an antiviral may be very effective when used early but may not be effective in late stage disease, and may even be harmful. Oseltamivir, for example, is generally only considered effective for influenza when used within 0-36 or 0-48 hours McLean, Treanor. Baloxavir studies for influenza also show that treatment delay is critical — Ikematsu report an 86% reduction in cases for post-exposure prophylaxis, Hayden show a 33 hour reduction in the time to alleviation of symptoms for treatment within 24 hours and a reduction of 13 hours for treatment within 24-48 hours, and Kumar report only 2.5 hours improvement for inpatient treatment.

Treatment delay	Result
Post exposure prophylaxis	86% fewer cases Ikematsu
<24 hours	-33 hours symptoms ^{Hayden}
24-48 hours	-13 hours symptoms Hayden
Inpatients	-2.5 hours to improvement Kumar

Table 3. Studies of baloxavir for influenza show that early treatment is more effective.

Figure 19 shows a mixed-effects meta-regression for efficacy as a function of treatment delay in COVID-19 studies from 66 treatments, showing that efficacy declines rapidly with treatment delay. Early treatment is critical for COVID-19.

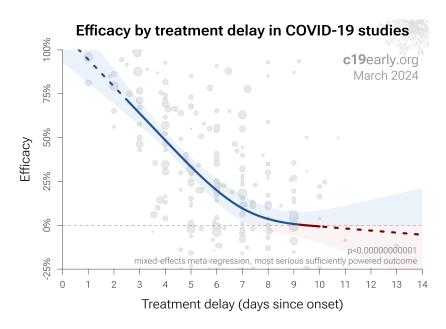


Figure 19. Early treatment is more effective. Meta-regression showing efficacy as a function of treatment delay in COVID-19 studies from 66 treatments.

Patient demographics. Details of the patient population including age and comorbidities may critically affect how well a treatment works. For example, many COVID-19 studies with relatively young low-comorbidity patients show all patients recovering quickly with or without treatment. In such cases, there is little room for an effective treatment to improve results (as in *López-Medina*).

Effect measured. Efficacy may differ significantly depending on the effect measured, for example a treatment may be very effective at reducing mortality, but less effective at minimizing cases or hospitalization. Or a treatment may have no effect on viral clearance while still being effective at reducing mortality.

Variants. There are many different variants of SARS-CoV-2 and efficacy may depend critically on the distribution of variants encountered by the patients in a study. For example, the Gamma variant shows significantly different characteristics *Faria, Karita, Nonaka, Zavascki*. Different mechanisms of action may be more or less effective depending on variants, for example the viral entry process for the omicron variant has moved towards TMPRSS2-independent fusion, suggesting that TMPRSS2 inhibitors may be less effective *Peacock, Willett*.

Regimen. Effectiveness may depend strongly on the dosage and treatment regimen.

Other treatments. The use of other treatments may significantly affect outcomes, including anything from supplements, other medications, or other kinds of treatment such as prone positioning.

Medication quality. The quality of medications may vary significantly between manufacturers and production batches, which may significantly affect efficacy and safety. *Williams* analyze ivermectin from 11 different sources, showing highly variable antiparasitic efficacy across different manufacturers. *Xu* analyze a treatment from two different manufacturers, showing 9 different impurities, with significantly different concentrations for each manufacturer. Non-prescription supplements may show very wide variations in quality *Crawford*, *Crighton*.

Pooled outcome analysis. We present both pooled analyses and specific outcome analyses. Notably, pooled analysis often results in earlier detection of efficacy as shown in Figure 20. For many COVID-19 treatments, a reduction in mortality logically follows from a reduction in hospitalization, which follows from a reduction in symptomatic cases, etc. An antiviral tested with a low-risk population may report zero mortality in both arms, however a reduction in severity and improved viral clearance may translate into lower mortality among a high-risk population, and including these results in pooled analysis allows faster detection of efficacy. Trials with high-risk patients may also be restricted due to ethical concerns for treatments that are known or expected to be effective.

Pooled analysis enables using more of the available information. While there is much more information available, for example dose-response relationships, the advantage of the method used here is simplicity and transparency. Note that pooled analysis could hide efficacy, for example a treatment that is beneficial for late stage patients but has no effect on viral replication or early stage disease could show no efficacy in pooled analysis if most studies only examine viral clearance. While we present pooled results, we also present individual outcome analyses, which may be more informative for specific use cases.

Pooled outcomes identify efficacy faster. Currently, 44 of the treatments we analyze show statistically significant efficacy or harm, defined as \geq 10% decreased risk or >0% increased risk from \geq 3 studies. 88% of treatments showing statistically significant efficacy/harm with pooled effects have been confirmed with one or more specific outcomes, with a mean delay of 3.6 months. When restricting to RCTs only, 50% of treatments showing statistically significant efficacy/harm with pooled effects have been confirmed with one or more specific outcomes, with a mean delay of 6.1 months.

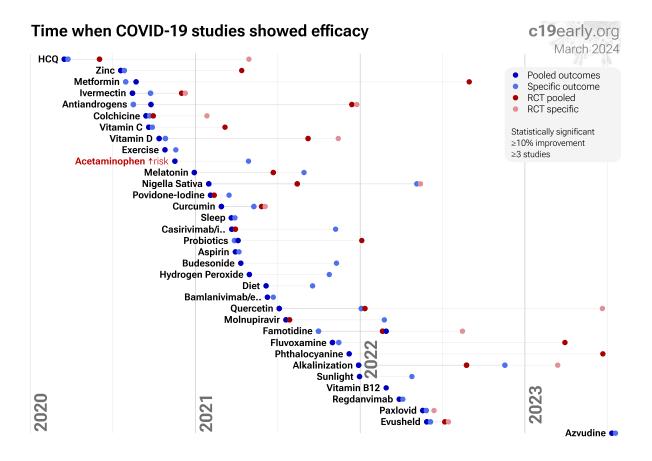


Figure 20. The time when studies showed that treatments were effective, defined as statistically significant improvement of ≥10% from ≥3 studies. Pooled results typically show efficacy earlier than specific outcome results. Results from all studies often shows efficacy much earlier than when restricting to RCTs. Results reflect conditions as used in trials to date, these depend on the population treated, treatment delay, and treatment regimen.

Meta analysis. The distribution of studies will alter the outcome of a meta analysis. Consider a simplified example where everything is equal except for the treatment delay, and effectiveness decreases to zero or below with increasing delay. If there are many studies using very late treatment, the outcome may be negative, even though early treatment is very effective. This may have a greater effect than pooling different outcomes such as mortality and hospitalization. For example a treatment may have 50% efficacy for mortality but only 40% for hospitalization when used within 48 hours. However efficacy could be 0% when used late.

All meta analyses combine heterogeneous studies, varying in population, variants, and potentially all factors above, and therefore may obscure efficacy by including studies where treatment is less effective. Generally, we expect the estimated effect size from meta analysis to be less than that for the optimal case. Looking at all studies is valuable for providing an overview of all research, important to avoid cherry-picking, and informative when a positive result is found despite combining less-optimal situations. However, the resulting estimate does not apply to specific cases such as early treatment in high-risk populations. While we present results for all studies, we also present treatment time and individual outcome analyses, which may be more informative for specific use cases.

Discussion

Publication bias. Publishing is often biased towards positive results, however evidence suggests that there may be a negative bias for inexpensive treatments for COVID-19. Both negative and positive results are very important for COVID-19, media in many countries prioritizes negative results for inexpensive treatments (inverting the typical

incentive for scientists that value media recognition), and there are many reports of difficulty publishing positive results *Boulware, Meeus, Meneguesso*. For vitamin A, there is currently not enough data to evaluate publication bias with high confidence.

One method to evaluate bias is to compare prospective vs. retrospective studies. Prospective studies are more likely to be published regardless of the result, while retrospective studies are more likely to exhibit bias. For example, researchers may perform preliminary analysis with minimal effort and the results may influence their decision to continue. Retrospective studies also provide more opportunities for the specifics of data extraction and adjustments to influence results.

Figure 21 shows a scatter plot of results for prospective and retrospective treatment studies. 50% of retrospective studies report a statistically significant positive effect for one or more outcomes, compared to 56% of prospective studies, consistent with a bias toward publishing negative results. The median effect size for retrospective studies is 19% improvement, compared to 64% for prospective studies, suggesting a potential bias towards publishing results showing lower efficacy.

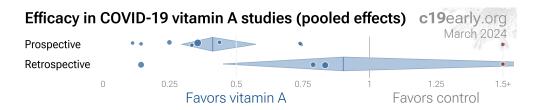


Figure 21. Prospective vs. retrospective studies. The diamonds show the results of random effects meta-analysis.

Funnel plot analysis. Funnel plots have traditionally been used for analyzing publication bias. This is invalid for COVID-19 acute treatment trials — the underlying assumptions are invalid, which we can demonstrate with a simple example. Consider a set of hypothetical perfect trials with no bias. Figure 22 plot A shows a funnel plot for a simulation of 80 perfect trials, with random group sizes, and each patient's outcome randomly sampled (10% control event probability, and a 30% effect size for treatment). Analysis shows no asymmetry (p > 0.05). In plot B, we add a single typical variation in COVID-19 treatment trials — treatment delay. Consider that efficacy varies from 90% for treatment within 24 hours, reducing to 10% when treatment is delayed 3 days. In plot B, each trial's treatment delay is randomly selected. Analysis now shows highly significant asymmetry, p < 0.0001, with six variants of Egger's test all showing p < 0.05 Egger, Harbord, Macaskill, Moreno, Peters, Rothstein, Rücker, Stanley. Note that these tests fail even though treatment delay is uniformly distributed. In reality treatment delay is more complex — each trial has a different distribution of delays across patients, and the distribution across trials may be biased (e.g., late treatment trials may be more common). Similarly, many other variations in trials may produce asymmetry, including dose, administration, duration of treatment, differences in SOC, comorbidities, age, variants, and bias in design, implementation, analysis, and reporting.

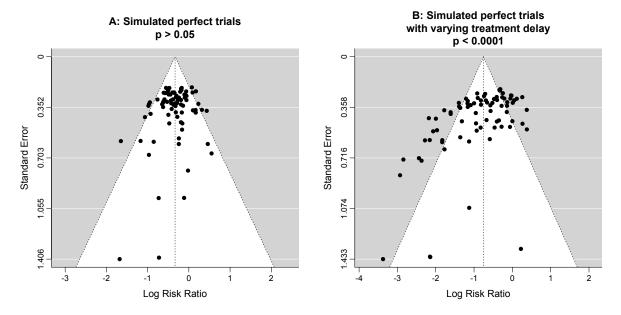


Figure 22. Example funnel plot analysis for simulated perfect trials.

Conflicts of interest. Pharmaceutical drug trials often have conflicts of interest whereby sponsors or trial staff have a financial interest in the outcome being positive. Vitamin A for COVID-19 lacks this because it is an inexpensive and widely available supplement. In contrast, most COVID-19 vitamin A trials have been run by physicians on the front lines with the primary goal of finding the best methods to save human lives and minimize the collateral damage caused by COVID-19. While pharmaceutical companies are careful to run trials under optimal conditions (for example, restricting patients to those most likely to benefit, only including patients that can be treated soon after onset when necessary, and ensuring accurate dosing), not all vitamin A trials represent the optimal conditions for efficacy.

Limitations. Summary statistics from meta analysis necessarily lose information. As with all meta analyses, studies are heterogeneous, with differences in treatment delay, treatment regimen, patient demographics, variants, conflicts of interest, standard of care, and other factors. We provide analyses by specific outcomes and by treatment delay, and we aim to identify key characteristics in the forest plots and summaries. Results should be viewed in the context of study characteristics.

Some analyses classify treatment based on early or late administration, as done here, while others distinguish between mild, moderate, and severe cases. Viral load does not indicate degree of symptoms — for example patients may have a high viral load while being asymptomatic. With regard to treatments that have antiviral properties, timing of treatment is critical — late administration may be less helpful regardless of severity.

Details of treatment delay per patient is often not available. For example, a study may treat 90% of patients relatively early, but the events driving the outcome may come from 10% of patients treated very late. Our 5 day cutoff for early treatment may be too conservative, 5 days may be too late in many cases.

Comparison across treatments is confounded by differences in the studies performed, for example dose, variants, and conflicts of interest. Trials affiliated with special interests may use designs better suited to the preferred outcome.

In some cases, the most serious outcome has very few events, resulting in lower confidence results being used in pooled analysis, however the method is simpler and more transparent. This is less critical as the number of studies increases. Restriction to outcomes with sufficient power may be beneficial in pooled analysis and improve accuracy when there are few studies, however we maintain our pre-specified method to avoid any retrospective changes.

Studies show that combinations of treatments can be highly synergistic and may result in many times greater efficacy than individual treatments alone Alsaidi, Andreani, De Forni, Fiaschi, Jeffreys, Jitobaom, Jitobaom (B), Ostrov, Said, Thairu, Wan. Therefore standard of care may be critical and benefits may diminish or disappear if standard of care does not include certain treatments.

This real-time analysis is constantly updated based on submissions. Accuracy benefits from widespread review and submission of updates and corrections from reviewers. Less popular treatments may receive fewer reviews.

No treatment, vaccine, or intervention is 100% available and effective for all current and future variants. Efficacy may vary significantly with different variants and within different populations. All treatments have potential side effects. Propensity to experience side effects may be predicted in advance by qualified physicians. We do not provide medical advice. Before taking any medication, consult a qualified physician who can compare all options, provide personalized advice, and provide details of risks and benefits based on individual medical history and situations.

Notes. 1 of 13 studies combine treatments. The results of vitamin A alone may differ. 1 of 5 RCTs use combined treatment.

Reviews. Multiple reviews cover vitamin A for COVID-19, presenting additional background on mechanisms and related results, including Andrade, DiGuilio (B), Midha, Stephensen.

Conclusion

Statistically significant lower risk is seen for recovery and viral clearance. 7 studies from 6 independent teams in 4 countries show statistically significant improvements. Meta analysis using the most serious outcome reported shows 38% [4-60%] lower risk. Results are similar for Randomized Controlled Trials, higher quality studies, and peer-reviewed studies. Results are consistent with early treatment being more effective than late treatment. 5 sufficiency studies analyze outcomes based on serum levels, showing 85% [54-95%] lower risk for patients with higher vitamin A levels. In exclusion sensitivity analysis, statistical significance is lost after excluding 2 of 13 studies in pooled analysis.

The European Food Safety Authority has found evidence for a causal relationship between the intake of vitamin A and optimal immune system function ^{Galmés}, ^{Galmés} (B).

Study Notes

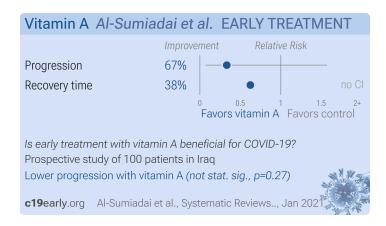
Al-Sumiadai

Al-Sumiadai (C): Treatment and prophylaxis studies of vitamin A in Iraq.

The treatment study contained 100 patients, 50 treated with 200,000IU vitamin A for two days, showing lower progression to severe disease, and shorter duration of symptoms.

The prophylaxis study contained 209 contacts of COVID-19 patients, 97 treated with vitamin A, showing significantly lower cases with treatment, and shorter duration of symptoms.

Al-Sumiadai

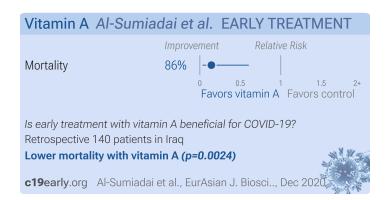


Al-Sumiadai (B): Treatment and prophylaxis studies of vitamin A in Iraq.

The treatment study contained 100 patients, 50 treated with 200,000IU vitamin A for two days, showing lower progression to severe disease, and shorter duration of symptoms.

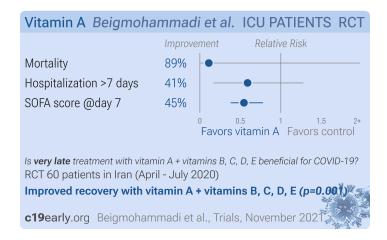
The prophylaxis study contained 209 contacts of COVID-19 patients, 97 treated with vitamin A, showing significantly lower cases with treatment, and shorter duration of symptoms.

Al-Sumiadai



Al-Sumiadai: Retrospective 70 severe condition patients treated with vitamin A (200,000IU for two days), salbutamol, and budesonide, and 70 patients not treated with vitamin A, showing significantly lower mortality with the addition of vitamin A.

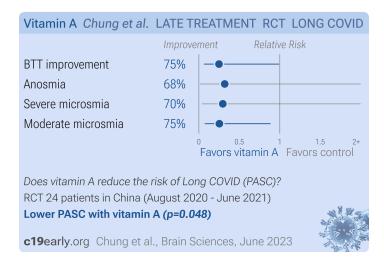
Beigmohammadi



Beigmohammadi: Small RCT 60 ICU patients in Iran, 30 treated with vitamins A, B, C, D, and E, showing significant improvement in SOFA score and several inflammatory markers at day 7 with treatment.

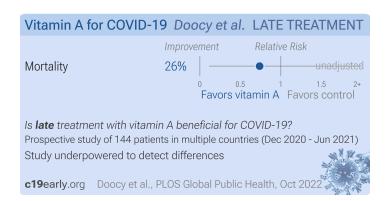
5,000 IU vitamin A daily, 600,000 IU vitamin D once, 300 IU of vitamin E twice a day, 500 mg vitamin C four times a day, and one ampule daily of B vitamins [thiamine nitrate 3.1 mg, sodium riboflavin phosphate 4.9 mg (corresponding to vitamin B2 3.6 mg), nicotinamide 40 mg, pyridoxine hydrochloride 4.9 mg (corresponding to vitamin B6 4.0 mg), sodium pantothenate 16.5 mg (corresponding to pantothenic acid 15 mg), sodium ascorbate 113 mg (corresponding to vitamin C 100 mg), biotin 60 μg, folic acid 400 μg, and cyanocobalamin 5 μg]. irct.ir.

Chung



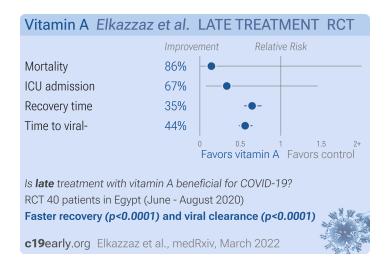
Chung: RCT 24 patients with olfactory dysfunction post-COVID-19 in Hong Kong, showing significantly improved recovery with the addition of vitamin A to aerosolised diffuser olfactory training. 25,000IU vitamin A for 14 days.

Doocy



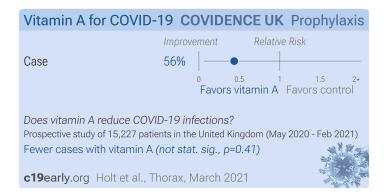
Doocy: Prospective study of 144 hospitalized COVID-19 patients in the DRC and South Sudan, showing no significant difference with vitamin A treatment in unadjusted results with only 8 patients receiving vitamin A.

Elkazzaz



Elkazzaz: RCT with 20 13-cis-retinoic acid patients and 20 control patients, showing faster recovery and viral clearance with treatment. Aerosolized 13-cis-retinoic acid with increasing dose from 0.2 mg/kg/day to 4 mg/kg/day for 14 days, plus oral 13-cis-retinoic acid 20 mg/day. 13-cis retinoic acid, also known as isotretinoin, is a synthetic vitamin A derivative that has been shown to have antiandrogenic effects.

Holt



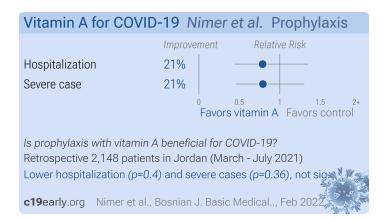
Holt: Prospective survey-based study with 15,227 people in the UK, showing lower risk of COVID-19 cases with vitamin A, vitamin D, zinc, selenium, probiotics, and inhaled corticosteroids; and higher risk with metformin and vitamin C. Statistical significance was not reached for any of these. Except for vitamin D, the results for treatments we follow were only adjusted for age, sex, duration of participation, and test frequency. NCT04330599. COVIDENCE UK.

Mandour



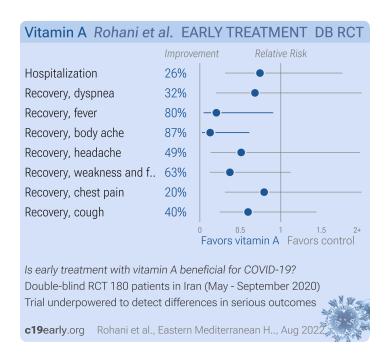
Mandour: Case control study with 30 ICU COVID-19 patients, 30 hospitalized non-ICU patients, and 30 matched healthy controls, showing vitamin A levels associated with COVID-19 and severity, with ICU patient levels < hospitalized patients < healthy controls. Authors also show significantly lower risk of ARDS with vitamin A levels above $0.65\mu g/ml$.

Nimer



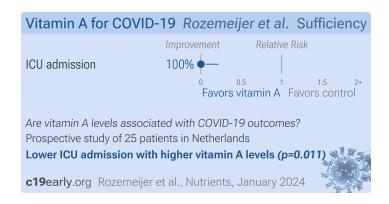
Nimer: Retrospective 2,148 COVID-19 recovered patients in Jordan, showing no significant differences in the risk of severity and hospitalization with vitamin A prophylaxis.

Rohani



Rohani: RCT 91 vitamin A and 91 control patients in Iran, showing improved recovery with treatment. All patients received HCQ. 25,000IU/day oral vitamin A for 10 days.

Rozemeijer



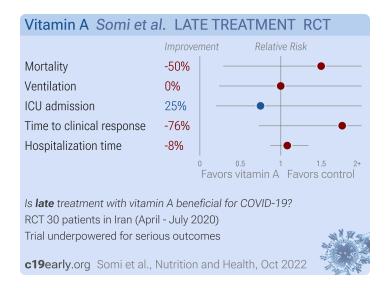
Rozemeijer: Prospective pilot study of 20 critically ill COVID-19 ICU patients showing high deficiency rates of 50-100% for vitamins A, B6, and D; zinc; and selenium at admission. Deficiencies of vitamins B6 and D, and low iron status, persisted after 3 weeks. Plasma levels of vitamins A and E, zinc, and selenium increased over time as inflammation resolved, suggesting redistribution may explain some observed deficiencies. All patients received daily micronutrient administration. Additional intravenous and oral micronutrient administration for 10 patients did not significantly impact micronutrient levels or deficiency rates, however authors note that the administered doses may be too low. The form of vitamin D is not specified but may have been cholecalciferol which is expected to have a very long onset of action compared to more appropriate forms such as calcifediol or calcitriol.

Sarohan



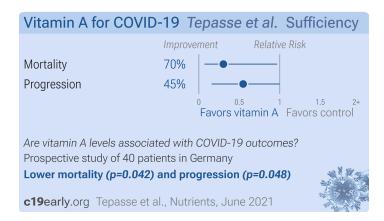
Sarohan: Retrospective 27 severe COVID-19 patients and 23 non-COVID-19 patients, showing significantly lower vitamin A levels in COVID-19 patients (0.37mg/L vs. 0.52 mg/L, p<0.001). 10 of 27 COVID-19 patients received vitamin A, with higher mortality. Group details are not provided but authors note that 8 of 10 had comorbidities.

Somi



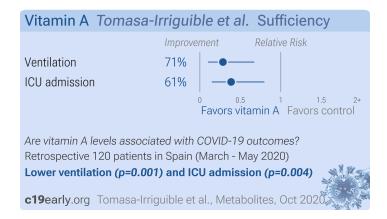
Somi: RCT 30 hospitalized patients in Iran, showing no significant difference with vitamin A treatment. All patients received HCQ. 50,000 IU/day intramuscular vitamin A for up to 2 weeks.

Tepasse



Tepasse: Prospective analysis of 40 hospitalized patients and 47 age-matched convalescent patients, showing significantly lower vitamin A levels in critical patients, and significantly lower vitamin A levels in hospitalized patients vs. controls. Low vitamin A levels were significantly associated with ARDS and mortality in hospitalized patients.

Tomasa-Irriguible

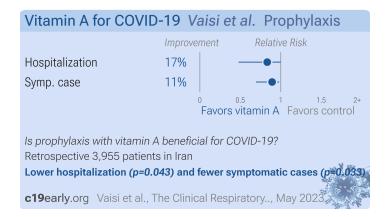


Tomasa-Irriguible (B): Retrospective 120 hospitalized patients in Spain showing vitamin A deficiency associated with higher ICU admission.

Tomasa-Irriguible

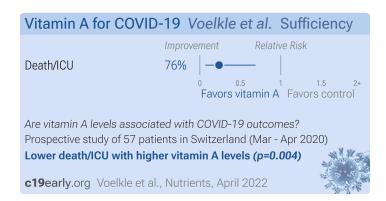
Tomasa-Irriguible: Estimated 300 patient vitamin A early treatment RCT with results expected soon (estimated completion over 3 months ago).

Vaisi



Vaisi: Analysis of nutrient intake and COVID-19 outcomes for 3,996 people in Iran, showing lower risk of COVID-19 hospitalization with sufficient vitamin A, vitamin C, and selenium intake, with statistical significance for vitamin A and selenium.

Voelkle



Voelkle: Prospective study of 57 consecutive hospitalized COVID-19 patients in Switzerland, showing higher risk of mortality/ICU admission with vitamin A, vitamin D, and zinc deficiency, with statistical significance only for vitamin A and zinc. Adjustments only considered age.

Appendix 1. Methods and Data

We perform ongoing searches of PubMed, medRxiv, Europe PMC, ClinicalTrials.gov, The Cochrane Library, Google Scholar, Research Square, ScienceDirect, Oxford University Press, the reference lists of other studies and meta-analyses, and submissions to the site c19early.org. Search terms are vitamin A and COVID-19 or SARS-CoV-2. Automated searches are performed twice daily, with all matches reviewed for inclusion. All studies regarding the use of vitamin A for COVID-19 that report a comparison with a control group are included in the main analysis. Sensitivity analysis is performed, excluding studies with major issues, epidemiological studies, and studies with minimal available information. This is a living analysis and is updated regularly.

We extracted effect sizes and associated data from all studies. If studies report multiple kinds of effects then the most serious outcome is used in pooled analysis, while other outcomes are included in the outcome specific analyses. For example, if effects for mortality and cases are both reported, the effect for mortality is used, this may be different to the effect that a study focused on. If symptomatic results are reported at multiple times, we used the latest time, for example if mortality results are provided at 14 days and 28 days, the results at 28 days have preference. Mortality alone is preferred over combined outcomes. Outcomes with zero events in both arms are not used, the next most serious outcome with one or more events is used. For example, in low-risk populations with no mortality, a reduction in mortality with treatment is not possible, however a reduction in hospitalization, for example, is still valuable. Clinical outcomes are considered more important than viral test status. When basically all patients recover in both treatment and control groups, preference for viral clearance and recovery is given to results mid-recovery where available. After most or all patients have recovered there is little or no room for an effective treatment to do better, however faster recovery is valuable. If only individual symptom data is available, the most serious symptom has priority, for example difficulty breathing or low SpO2 is more important than cough. When results provide an odds ratio, we compute the relative risk when possible, or convert to a relative risk according to Zhang. Reported confidence intervals and p-values were used when available, using adjusted values when provided. If multiple types of adjustments are reported propensity score matching and multivariable regression has preference over propensity score matching or weighting, which has preference over multivariable regression. Adjusted results have preference over unadjusted results for a more serious outcome when the adjustments significantly alter results. When needed, conversion between reported pvalues and confidence intervals followed Altman, Altman (B), and Fisher's exact test was used to calculate p-values for event data. If continuity correction for zero values is required, we use the reciprocal of the opposite arm with the sum of the correction factors equal to 1 Sweeting. Results are expressed with RR < 1.0 favoring treatment, and using the risk of a negative outcome when applicable (for example, the risk of death rather than the risk of survival). If studies only report relative continuous values such as relative times, the ratio of the time for the treatment group versus the time for the control group is used. Calculations are done in Python (3.12.2) with scipy (1.12.0), pythonmeta (1.26), numpy (1.26.4), statsmodels (0.14.1), and plotly (5.19.0).

Forest plots are computed using PythonMeta ^{Deng} with the DerSimonian and Laird random effects model (the fixed effect assumption is not plausible in this case) and inverse variance weighting. Results are presented with 95% confidence intervals. Heterogeneity among studies was assessed using the I² statistic. Mixed-effects meta-regression results are computed with R (4.1.2) using the metafor (3.0-2) and rms (6.2-0) packages, and using the most serious sufficiently powered outcome. For all statistical tests, a p-value less than 0.05 was considered statistically significant. Grobid 0.8.0 is used to parse PDF documents.

We have classified studies as early treatment if most patients are not already at a severe stage at the time of treatment (for example based on oxygen status or lung involvement), and treatment started within 5 days of the onset of symptoms. If studies contain a mix of early treatment and late treatment patients, we consider the treatment time of patients contributing most to the events (for example, consider a study where most patients are treated early but late treatment patients are included, and all mortality events were observed with late treatment patients). We note that a shorter time may be preferable. Antivirals are typically only considered effective when used within a shorter timeframe, for example 0-36 or 0-48 hours for oseltamivir, with longer delays not being effective McLean, Treanor.

We received no funding, this research is done in our spare time. We have no affiliations with any pharmaceutical companies or political parties.

A summary of study results is below. Please submit updates and corrections at https://c19early.org/vameta.html.

Early treatment

Effect extraction follows pre-specified rules as detailed above and gives priority to more serious outcomes. For pooled analyses, the first (most serious) outcome is used, which may differ from the effect a paper focuses on. Other outcomes are used in outcome specific analyses.

Al-Sumiadai (B), 1/31/2021, prospective, Iraq, preprint, 3 authors.

risk of progression, 66.7% lower, RR 0.33, p = 0.27, treatment 2 of 50 (4.0%), control 6 of 50 (12.0%), NNT 13, progression to severe disease.

Al-Sumiadai, 12/31/2020, retrospective, Iraq, peer-reviewed, 3 authors, excluded in exclusion analyses: minimal details of groups provided.	risk of death, 85.7% lower, RR 0.14, <i>p</i> = 0.002, treatment 2 of 70 (2.9%), control 14 of 70 (20.0%), NNT 5.8.
Rohani, 8/18/2022, Double Blind Randomized Controlled Trial, placebo-controlled, Iran, peerreviewed, mean age 39.4, 6 authors, study period 1 May, 2020 - 1 September, 2020, trial IRCT46974.	risk of hospitalization, 25.6% lower, RR 0.74, p = 0.63, treatment 8 of 89 (9.0%), control 11 of 91 (12.1%), NNT 32.
	risk of no recovery, 31.8% lower, RR 0.68, <i>p</i> = 0.53, treatment 4 of 89 (4.5%), control 6 of 91 (6.6%), NNT 48, dyspnea.
	risk of no recovery, 79.6% lower, RR 0.20, <i>p</i> = 0.03, treatment 2 of 89 (2.2%), control 10 of 91 (11.0%), NNT 11, fever.
	risk of no recovery, 87.2% lower, RR 0.13, <i>p</i> = 0.01, treatment 1 of 89 (1.1%), control 8 of 91 (8.8%), NNT 13, body ache.
	risk of no recovery, 48.9% lower, RR 0.51, <i>p</i> = 0.32, treatment 3 of 89 (3.4%), control 6 of 91 (6.6%), NNT 31, headache.
	risk of no recovery, 62.8% lower, RR 0.37, p = 0.05, treatment 4 of 89 (4.5%), control 11 of 91 (12.1%), NNT 13, weakness and fatigue.
	risk of no recovery, 20.5% lower, RR 0.80, <i>p</i> = 0.63, treatment 7 of 89 (7.9%), control 9 of 91 (9.9%), NNT 49, chest pain.
	risk of no recovery, 40.4% lower, RR 0.60, <i>p</i> = 0.24, treatment 7 of 89 (7.9%), control 12 of 91 (13.2%), NNT 19, cough.
Tomasa-Irriguible, 11/30/2023, Double Blind Randomized Controlled Trial, placebo-controlled, Spain, trial NCT04751669 (history) (CoVIT).	Estimated 300 patient RCT with results unknown and over 3 months late.

Late treatment

Effect extraction follows pre-specified rules as detailed above and gives priority to more serious outcomes. For pooled analyses, the first (most serious) outcome is used, which may differ from the effect a paper focuses on. Other outcomes are used in outcome specific analyses.

Beigmohammadi, 11/14/2021, Single Blind Randomized Controlled Trial, Iran, peer-reviewed, 6 authors, study period April 2020 - July 2020, this trial uses multiple treatments in the treatment arm (combined with vitamins B, C, D, E) - results of	risk of death, 88.9% lower, RR 0.11, p = 0.11, treatment 0 of 30 (0.0%), control 4 of 30 (13.3%), NNT 7.5, relative risk is not 0 because of continuity correction due to zero events (with reciprocal of the contrasting arm).	
individual treatments may vary, trial IRCT20200319046819N1.	risk of hospitalization >7 days, 41.0% lower, RR 0.59, p = 0.25, treatment 4 of 30 (13.3%), control 16 of 30 (53.3%), NNT 2.5, adjusted per study, odds ratio converted to relative risk.	
	relative SOFA score @day 7, 45.5% better, RR 0.55, p < 0.001, treatment 30, control 30.	
Chung, 6/30/2023, Randomized Controlled Trial, China, peer-reviewed, 14 authors, study period 14 August, 2020 - 11 June, 2021, trial NCT04900415 (history).	relative BTT improvement, 75.1% better, RR 0.25, p = 0.048, treatment mean 3.01 (±2.52) n=9, control mean 0.75 (±1.67) n=8, vitamin A + OT vs. OT.	

	anosmia, 68.0% lower, RR 0.32, p = 0.47, treatment 0 of 9 (0.0%), control 1 of 8 (12.5%), NNT 8.0, relative risk is not 0 because of continuity correction due to zero events (with reciprocal of the contrasting arm), vitamin A + OT vs. OT.
	severe microsmia, 70.4% lower, RR 0.30, p = 0.29, treatment 1 of 9 (11.1%), control 3 of 8 (37.5%), NNT 3.8, vitamin A + OT vs OT.
	moderate microsmia, 74.6% lower, RR 0.25, <i>p</i> = 0.02, treatmen 2 of 9 (22.2%), control 7 of 8 (87.5%), NNT 1.5, vitamin A + OT vs. OT.
Doocy, 10/19/2022, prospective, multiple countries, peer-reviewed, 6 authors, study period December 2020 - June 2021, trial NCT04568499 (history), excluded in exclusion analyses: unadjusted results with no group details.	risk of death, 26.1% lower, RR 0.74, <i>p</i> = 1.00, treatment 1 of 8 (12.5%), control 23 of 136 (16.9%), NNT 23, unadjusted.
Elkazzaz, 3/8/2022, Randomized Controlled Trial, Egypt, preprint, 4 authors, study period June 2020 - August 2020.	risk of death, 85.7% lower, RR 0.14, p = 0.23, treatment 0 of 20 (0.0%), control 3 of 20 (15.0%), NNT 6.7, relative risk is not 0 because of continuity correction due to zero events (with reciprocal of the contrasting arm).
	risk of ICU admission, 66.7% lower, RR 0.33, p = 0.24, treatmen 2 of 20 (10.0%), control 6 of 20 (30.0%), NNT 5.0.
	recovery time, 35.4% lower, relative time 0.65, p < 0.001, treatment mean 16.3 (±4.5) n=20, control mean 25.23 (±4.72) n=20.
	time to viral-, 44.0% lower, relative time 0.56, $p < 0.001$, treatment mean 13.36 (±1.49) n=20, control mean 23.85 (±4.0) n=20.
Sarohan, 2/1/2021, retrospective, Turkey, preprint, 4 authors, excluded in exclusion analyses: unadjusted results with no group details, comments suggest significant group differences and confounding.	risk of death, 282.5% higher, RR 3.83, <i>p</i> = 0.001, treatment 9 of 10 (90.0%), control 4 of 17 (23.5%).
Somi, 10/7/2022, Randomized Controlled Trial, Iran, peer-reviewed, mean age 60.2, 7 authors, study period April 2020 - July 2020, trial IRCT20170117032004N3.	risk of death, 50.0% higher, RR 1.50, <i>p</i> = 1.00, treatment 3 of 19 (20.0%), control 2 of 15 (13.3%).
	risk of mechanical ventilation, no change, RR 1.00, p = 1.00, treatment 3 of 15 (20.0%), control 3 of 15 (20.0%).
	risk of ICU admission, 25.0% lower, RR 0.75, <i>p</i> = 1.00, treatmen 3 of 15 (20.0%), control 4 of 15 (26.7%), NNT 15.
	time to clinical response, 76.0% higher, HR 1.76, $p = 0.21$, treatment 15, control 15, Kaplan–Meier.
	hospitalization time, 8.1% higher, relative time 1.08, p = 0.49, treatment mean 7.33 (±2.31) n=15, control mean 6.78 (±1.84) n=15.

Prophylaxis

Effect extraction follows pre-specified rules as detailed above and gives priority to more serious outcomes. For pooled analyses, the first (most serious) outcome is used, which may differ from the effect a paper focuses on. Other outcomes are used in outcome specific analyses.

Al-Sumiadai (C), 1/31/2021, prospective, Iraq, preprint, 3 authors.	risk of case, 64.5% lower, RR 0.36, <i>p</i> < 0.001, treatment 20 of 97 (20.6%), control 65 of 112 (58.0%), NNT 2.7.
Holt, 3/30/2021, prospective, United Kingdom, peer-reviewed, 34 authors, study period 1 May, 2020 - 5 February, 2021, trial NCT04330599 (history) (COVIDENCE UK), excluded in exclusion analyses: significant unadjusted confounding possible.	risk of case, 56.3% lower, RR 0.44, p = 0.41, treatment 1 of 91 (1.1%), control 445 of 15,136 (2.9%), NNT 54, adjusted per study, odds ratio converted to relative risk, minimally adjusted, group sizes approximated.
Nimer, 2/28/2022, retrospective, Jordan, peer- reviewed, survey, 4 authors, study period March 2021 - July 2021.	risk of hospitalization, 21.2% lower, RR 0.79, p = 0.40, treatment 15 of 144 (10.4%), control 204 of 2,004 (10.2%), adjusted per study, odds ratio converted to relative risk, multivariable.
	risk of severe case, 20.8% lower, RR 0.79, p = 0.36, treatment 17 of 144 (11.8%), control 243 of 2,004 (12.1%), adjusted per study, odds ratio converted to relative risk, multivariable.
Vaisi, 5/11/2023, retrospective, Iran, peer-reviewed, 5 authors.	risk of hospitalization, 16.7% lower, HR 0.83, p = 0.04, treatment 1,140, control 2,815, adjusted per study, inverted to make HR<1 favor treatment, sufficient vs. insufficient intake, multivariable, Cox proportional hazards.
	risk of symptomatic case, 10.6% lower, HR 0.89, p = 0.03, treatment 1,140, control 2,815, adjusted per study, inverted to make HR<1 favor treatment, sufficient vs. insufficient intake, multivariable, Cox proportional hazards.

Supplementary Data

Supplementary Data

Footnotes

a. Viral infection and replication involves attachment, entry, uncoating and release, genome replication and transcription, translation and protein processing, assembly and budding, and release. Each step can be disrupted by therapeutics.

References

- 1. **Al-Sumiadai** et al., *Therapeutic effect of Vitamin A on severe COVID-19 patients*, EurAsian Journal of Biosciences, 14:7347-7350, ejobios.org/article/therapeutic-effect-of-vitamin-a-on-severe-covid-19-patients-8517.
- Al-Sumiadai (B) et al., Therapeutic effect of vitamin A on COVID-19 patients and its prophylactic effect on contacts, Systematic Reviews in Pharmacy, 12:1, www.researchgate.net/publication/351637178_THERAPEUTIC_EFFECT_OF_VITAMIN_A_ON_COVID-19_PATIENTS_AND_ITS_PROPH YLACTIC_EFFECT_ON_CONTACTS.
- 3. **Al-Sumiadai (C)** et al., Therapeutic effect of vitamin A on COVID-19 patients and its prophylactic effect on contacts, Systematic Reviews in Pharmacy, 12:1, www.researchgate.net/publication/351637178_THERAPEUTIC_EFFECT_OF_VITAMIN_A_ON_COVID-19_PATIENTS_AND_ITS_PROPH YLACTIC_EFFECT_ON_CONTACTS.
- 4. Als-Nielsen et al., Association of Funding and Conclusions in Randomized Drug Trials, JAMA, doi:10.1001/jama.290.7.921.
- 5. **Alsaidi** et al., Griffithsin and Carrageenan Combination Results in Antiviral Synergy against SARS-CoV-1 and 2 in a Pseudoviral Model, Marine Drugs, doi:10.3390/md19080418.
- 6. Altman, D., How to obtain the P value from a confidence interval, BMJ, doi:10.1136/bmj.d2304.
- 7. Altman (B) et al., How to obtain the confidence interval from a P value, BMJ, doi:10.1136/bmj.d2090.
- 8. **Andrade** et al., Vitamin A and D deficiencies in the prognosis of respiratory tract infections: A systematic review with perspectives for COVID-19 and a critical analysis on supplementation, SciELO preprints, doi:10.1590/SciELOPreprints.839.
- 9. **Andreani** et al., *In vitro* testing of combined hydroxychloroquine and azithromycin on SARS-CoV-2 shows synergistic effect, Microbial Pathogenesis, doi:/10.1016/j.micpath.2020.104228.
- 10. **Anglemyer** et al., Healthcare outcomes assessed with observational study designs compared with those assessed in randomized trials, Cochrane Database of Systematic Reviews 2014, Issue 4, doi:10.1002/14651858.MR000034.pub2.
- 11. **Beigmohammadi** et al., The effect of supplementation with vitamins A, B, C, D, and E on disease severity and inflammatory responses in patients with COVID-19: a randomized clinical trial, Trials, doi:10.1186/s13063-021-05795-4.
- 12. Boulware, D., Comments regarding paper rejection, twitter.com/boulware_dr/status/1311331372884205570.
- 13. c19early.org, c19early.org/timeline.html.
- 14. c19early.org (B), c19early.org/treatments.html.
- 15. **Chakraborty** et al., *In-silico* screening and in-vitro assay show the antiviral effect of Indomethacin against SARS-CoV-2, Computers in Biology and Medicine, doi:10.1016/j.compbiomed.2022.105788.
- 16. **Chung** et al., A Pilot Study of Short-Course Oral Vitamin A and Aerosolised Diffuser Olfactory Training for the Treatment of Smell Loss in Long COVID, Brain Sciences, doi:10.3390/brainsci13071014.
- 17. Concato et al., NEJM, 342:1887-1892, doi:10.1056/NEJM200006223422507.
- 18. **Crawford** et al., Analysis of Select Dietary Supplement Products Marketed to Support or Boost the Immune System, JAMA Network Open, doi:10.1001/jamanetworkopen.2022.26040.
- 19. **Crighton** et al., Toxicological screening and DNA sequencing detects contamination and adulteration in regulated herbal medicines and supplements for diet, weight loss and cardiovascular health, Journal of Pharmaceutical and Biomedical Analysis, doi:10.1016/j.jpba.2019.112834.
- 20. **Davidson** et al., No evidence of important difference in summary treatment effects between COVID-19 preprints and peer-reviewed publications: a meta-epidemiological study, Journal of Clinical Epidemiology, doi:10.1016/j.jclinepi.2023.08.011.
- 21. **De Forni** et al., Synergistic drug combinations designed to fully suppress SARS-CoV-2 in the lung of COVID-19 patients, PLoS ONE, doi:10.1371/journal.pone.0276751.

- 22. **Deaton** et al., *Understanding and misunderstanding randomized controlled trials*, Social Science & Medicine, 210, doi:10.1016/j.socscimed.2017.12.005.
- 23. Deng, H., PyMeta, Python module for meta-analysis, www.pymeta.com/.
- 24. **DiGuilio** et al., The multiphasic TNF-α-induced compromise of Calu-3 airway epithelial barrier function, Experimental Lung Research, doi:10.1080/01902148.2023.2193637.
- 25. **DiGuilio (B)** et al., Micronutrient Improvement of Epithelial Barrier Function in Various Disease States: A Case for Adjuvant Therapy, International Journal of Molecular Sciences, doi:10.3390/ijms23062995.
- 26. **Doocy** et al., Clinical progression and outcomes of patients hospitalized with COVID-19 in humanitarian settings: A prospective cohort study in South Sudan and Eastern Democratic Republic of the Congo, PLOS Global Public Health, doi:10.1371/journal.pgph.0000924.
- 27. **Eberhardt** et al., SARS-CoV-2 infection triggers pro-atherogenic inflammatory responses in human coronary vessels, Nature Cardiovascular Research, doi:10.1038/s44161-023-00336-5.
- 28. **EFSA**, Scientific Opinion on the substantiation of health claims related to vitamin A and cell differentiation (ID 14), function of the immune system (ID 14), maintenance of skin and mucous membranes (ID 15, 17), maintenance of vision (ID 16), maintenance of bone (ID 13, 17), maintenance of teeth (ID 13, 17), maintenance of hair (ID 17), metabolism of iron (ID 206), and protection of DNA, proteins and lipids from oxidative damage (ID 209) pursuant to Article 13(1) of Regulation (EC) No 1924/2006, EFSA Journal, doi:10.2903/j.efsa.2009.1221.
- 29. Egger et al., Bias in meta-analysis detected by a simple, graphical test, BMJ, doi:10.1136/bmj.315.7109.629.
- 30. **Elkazzaz** et al., 13 cis retinoic acid improved the outcomes of COVID-19 patients. A randomized clinical trial, medRxiv, doi:10.1101/2022.03.05.22271959.
- 31. Faria et al., Genomics and epidemiology of the P.1 SARS-CoV-2 lineage in Manaus, Brazil, Science, doi:10.1126/science.abh2644.
- 32. **Fiaschi** et al., In Vitro Combinatorial Activity of Direct Acting Antivirals and Monoclonal Antibodies against the Ancestral B.1 and BQ.1.1 SARS-CoV-2 Viral Variants, Viruses, doi:10.3390/v16020168.
- 33. **Franco** et al., Retinoic Acid-Mediated Inhibition of Mouse Coronavirus Replication Is Dependent on IRF3 and CaMKK, Viruses, doi:10.3390/v16010140.
- 34. **Galmés** et al., Suboptimal Consumption of Relevant Immune System Micronutrients Is Associated with a Worse Impact of COVID-19 in Spanish Populations, Nutrients, doi:10.3390/nu14112254.
- 35. **Galmés (B)** et al., Current State of Evidence: Influence of Nutritional and Nutrigenetic Factors on Immunity in the COVID-19 Pandemic Framework, Nutrients, doi:10.3390/nu12092738.
- 36. **Gøtzsche**, P., *Bias in double-blind trials*, Doctoral Thesis, University of Copenhagen, www.scientificfreedom.dk/2023/05/16/bias-in-double-blind-trials-doctoral-thesis/.
- 37. **Harbord** et al., A modified test for small-study effects in meta-analyses of controlled trials with binary endpoints, Statistics in Medicine, doi:10.1002/sim.2380.
- 38. **Hayden** et al., *Baloxavir Marboxil for Uncomplicated Influenza in Adults and Adolescents*, New England Journal of Medicine, doi:10.1056/NEJMoa1716197.
- 39. **Holt** et al., *Risk factors for developing COVID-19: a population-based longitudinal study (COVIDENCE UK)*, Thorax, doi:10.1136/thoraxjnl-2021-217487.
- 40. **Huang** et al., All-trans retinoic acid acts as a dual-purpose inhibitor of SARS-CoV-2 infection and inflammation, Computers in Biology and Medicine, doi:10.1016/j.compbiomed.2024.107942.
- 41. **Ikematsu** et al., Baloxavir Marboxil for Prophylaxis against Influenza in Household Contacts, New England Journal of Medicine, doi:10.1056/NEJMoa1915341.
- 42. irct.ir, www.irct.ir/trial/46838.
- 43. Jadad et al., Randomized Controlled Trials: Questions, Answers, and Musings, Second Edition, doi:10.1002/9780470691922.

- 44. **Jeffreys** et al., Remdesivir-ivermectin combination displays synergistic interaction with improved in vitro activity against SARS-CoV-2, International Journal of Antimicrobial Agents, doi:10.1016/j.ijantimicag.2022.106542.
- 45. **Jitobaom** et al., Favipiravir and Ivermectin Showed in Vitro Synergistic Antiviral Activity against SARS-CoV-2, Research Square, doi:10.21203/rs.3.rs-941811/v1.
- 46. **Jitobaom (B)** et al., Synergistic anti-SARS-CoV-2 activity of repurposed anti-parasitic drug combinations, BMC Pharmacology and Toxicology, doi:10.1186/s40360-022-00580-8.
- 47. **Karita** et al., Trajectory of viral load in a prospective population-based cohort with incident SARS-CoV-2 G614 infection, medRxiv, doi:10.1101/2021.08.27.21262754.
- 48. **Kumar** et al., Combining baloxavir marboxil with standard-of-care neuraminidase inhibitor in patients hospitalised with severe influenza (FLAGSTONE): a randomised, parallel-group, double-blind, placebo-controlled, superiority trial, The Lancet Infectious Diseases, doi:10.1016/S1473-3099(21)00469-2.
- 49. **Lee** et al., *Analysis of Overall Level of Evidence Behind Infectious Diseases Society of America Practice Guidelines*, Arch Intern Med., 2011, 171:1, 18-22, doi:10.1001/archinternmed.2010.482.
- 50. Li et al., Revealing the targets and mechanisms of vitamin A in the treatment of COVID-19, Aging, doi:10.18632/aging.103888.
- 51. **López-Medina** et al., Effect of Ivermectin on Time to Resolution of Symptoms Among Adults With Mild COVID-19: A Randomized Clinical Trial, JAMA, doi:10.1001/jama.2021.3071.
- 52. Lui et al., Nsp1 facilitates SARS-CoV-2 replication through calcineurin-NFAT signaling, Virology, doi:10.1128/mbio.00392-24.
- 53. Lv et al., Host proviral and antiviral factors for SARS-CoV-2, Virus Genes, doi:10.1007/s11262-021-01869-2.
- 54. **Macaskill** et al., A comparison of methods to detect publication bias in meta-analysis, Statistics in Medicine, doi:10.1002/sim.698.
- 55. **Malone** et al., Structures and functions of coronavirus replication–transcription complexes and their relevance for SARS-CoV-2 drug design, Nature Reviews Molecular Cell Biology, doi:10.1038/s41580-021-00432-z.
- 56. **Mandour** et al., Evaluation of vitamin A and E deficiency with severity of SARS-COV-2 disease: a case—control study, The Egyptian Journal of Bronchology, doi:10.1186/s43168-023-00210-9.
- 57. **McLean** et al., Impact of Late Oseltamivir Treatment on Influenza Symptoms in the Outpatient Setting: Results of a Randomized Trial, Open Forum Infect. Dis. September 2015, 2:3, doi:10.1093/ofid/ofv100.
- 58. **Meeus**, G., *Online Comment*, twitter.com/gertmeeus_MD/status/1386636373889781761.
- 59. Meneguesso, A., Médica defende tratamento precoce da Covid-19, www.youtube.com/watch?v=X5FCrlm_19U.
- 60. **Midha** et al., Mega doses of retinol: A possible immunomodulation in Covid-19 illness in resource-limited settings, Reviews in Medical Virology, doi:10.1002/rmv.2204.
- 61. **Moatasim** et al., Potent Antiviral Activity of Vitamin B12 against Severe Acute Respiratory Syndrome Coronavirus 2, Middle East Respiratory Syndrome Coronavirus, and Human Coronavirus 229E, Microorganisms, doi:10.3390/microorganisms11112777.
- 62. **Moreno** et al., Assessment of regression-based methods to adjust for publication bias through a comprehensive simulation study, BMC Medical Research Methodology, doi:10.1186/1471-2288-9-2.
- 63. **Morita** et al., All-Trans Retinoic Acid Exhibits Antiviral Effect against SARS-CoV-2 by Inhibiting 3CLpro Activity, Viruses, doi:10.3390/v13081669.
- 64. **Murigneux** et al., Proteomic analysis of SARS-CoV-2 particles unveils a key role of G3BP proteins in viral assembly, Nature Communications, doi:10.1038/s41467-024-44958-0.
- 65. **Nichol** et al., *Challenging issues in randomised controlled trials*, Injury, 2010, doi: 10.1016/j.injury.2010.03.033, www.injuryjournal.com/article/S0020-1383(10)00233-0/fulltext.
- 66. **Nimer** et al., The impact of vitamin and mineral supplements usage prior to COVID-19 infection on disease severity and hospitalization, Bosnian Journal of Basic Medical Sciences, doi:10.17305/bjbms.2021.7009.

- 67. **Nonaka** et al., SARS-CoV-2 variant of concern P.1 (Gamma) infection in young and middle-aged patients admitted to the intensive care units of a single hospital in Salvador, Northeast Brazil, February 2021, International Journal of Infectious Diseases, doi:10.1016/j.ijid.2021.08.003.
- 68. **Ostrov** et al., Highly Specific Sigma Receptor Ligands Exhibit Anti-Viral Properties in SARS-CoV-2 Infected Cells, Pathogens, doi:10.3390/pathogens10111514.
- 69. **Pandya** et al., *Unravelling Vitamin B12 as a potential inhibitor against SARS-CoV-2: A computational approach*, Informatics in Medicine Unlocked, doi:10.1016/j.imu.2022.100951.
- 70. **Peacock** et al., The SARS-CoV-2 variant, Omicron, shows rapid replication in human primary nasal epithelial cultures and efficiently uses the endosomal route of entry, bioRxiv, doi:10.1101/2021.12.31.474653.
- 71. Peters, J., Comparison of Two Methods to Detect Publication Bias in Meta-analysis, JAMA, doi:10.1001/jama.295.6.676.
- 72. **Rohani** et al., Evaluation and comparison of the effect of vitamin A supplementation with standard therapies in the treatment of patients with COVID-19, Eastern Mediterranean Health Journal, doi:10.26719/emhj.22.064.
- 73. **Rothstein**, H., *Publication Bias in Meta-Analysis: Prevention, Assessment and Adjustments*, www.wiley.com/en-ae/Publication+Bias+in+Meta+Analysis:+Prevention,+Assessment+and+Adjustments-p-9780470870143.
- 74. Rozemeijer et al., Micronutrient Status of Critically III Patients with COVID-19 Pneumonia, Nutrients, doi:10.3390/nu16030385.
- 75. **Rücker** et al., Arcsine test for publication bias in meta-analyses with binary outcomes, Statistics in Medicine, doi:10.1002/sim.2971.
- 76. **Said** et al., The effect of Nigella sativa and vitamin D3 supplementation on the clinical outcome in COVID-19 patients: A randomized controlled clinical trial, Frontiers in Pharmacology, doi:10.3389/fphar.2022.1011522.
- 77. Sarohan et al., Retinol Depletion in Severe COVID-19, medRxiv, doi:10.1101/2021.01.30.21250844.
- 78. **Scardua-Silva** et al., *Microstructural brain abnormalities, fatigue, and cognitive dysfunction after mild COVID-19*, Scientific Reports, doi:10.1038/s41598-024-52005-7.
- 79. **Somi** et al., Effect of vitamin A supplementation on the outcome severity of COVID-19 in hospitalized patients: A pilot randomized clinical trial, Nutrition and Health, doi:10.1177/02601060221129144.
- 80. **Stanley** et al., *Meta-regression approximations to reduce publication selection bias*, Research Synthesis Methods, doi:10.1002/jrsm.1095.
- 81. **Stephensen** et al., *Vitamin A in resistance to and recovery from infection: relevance to SARS-CoV2*, British Journal of Nutrition, doi:10.1017/S0007114521000246.
- 82. **Sweeting** et al., What to add to nothing? Use and avoidance of continuity corrections in meta-analysis of sparse data, Statistics in Medicine, doi:10.1002/sim.1761.
- 83. **Tepasse** et al., *Vitamin A Plasma Levels in COVID-19 Patients: A Prospective Multicenter Study and Hypothesis*, Nutrients, doi:10.3390/nu13072173.
- 84. **Thairu** et al., A Comparison of Ivermectin and Non Ivermectin Based Regimen for COVID-19 in Abuja: Effects on Virus Clearance, Days-to-discharge and Mortality, Journal of Pharmaceutical Research International, doi:10.9734/jpri/2022/v34i44A36328.
- 85. **Tomasa-Irriguible** et al., Efficacy of Micronutrient Dietary Supplementation in Reducing Hospital Admissions for COVID-19: A Double-blind, Placebo-controlled, Randomized Clinical Trial, NCT04751669, clinicaltrials.gov/study/NCT04751669.
- 86. **Tomasa-Irriguible (B)** et al., Low Levels of Few Micronutrients May Impact COVID-19 Disease Progression: An Observational Study on the First Wave, Metabolites, doi:10.3390/metabo11090565.
- 87. **Tong** et al., A Retinol Derivative Inhibits SARS-CoV-2 Infection by Interrupting Spike-Mediated Cellular Entry, mBio, doi:10.1128/mbio.01485-22.
- 88. **Treanor** et al., Efficacy and Safety of the Oral Neuraminidase Inhibitor Oseltamivir in Treating Acute Influenza: A Randomized Controlled Trial, JAMA, 2000, 283:8, 1016-1024, doi:10.1001/jama.283.8.1016.

- 89. **Vaisi** et al., The association between nutrients and occurrence of COVID-19 outcomes in the population of Western Iran: A cohort study, The Clinical Respiratory Journal, doi:10.1111/crj.13632.
- 90. **Voelkle** et al., Prevalence of Micronutrient Deficiencies in Patients Hospitalized with COVID-19: An Observational Cohort Study, Nutrients, doi:10.3390/nu14091862.
- 91. **Wan** et al., Synergistic inhibition effects of andrographolide and baicalin on coronavirus mechanisms by downregulation of ACE2 protein level, Scientific Reports, doi:10.1038/s41598-024-54722-5.
- 92. **Willett** et al., The hyper-transmissible SARS-CoV-2 Omicron variant exhibits significant antigenic change, vaccine escape and a switch in cell entry mechanism, medRxiv, doi:10.1101/2022.01.03.21268111.
- 93. **Williams**, T., Not All Ivermectin Is Created Equal: Comparing The Quality of 11 Different Ivermectin Sources, Do Your Own Research, doyourownresearch.substack.com/p/not-all-ivermectin-is-created-equal.
- 94. **Xu** et al., A study of impurities in the repurposed COVID-19 drug hydroxychloroquine sulfate by UHPLC-Q/TOF-MS and LC-SPE-NMR, Rapid Communications in Mass Spectrometry, doi:10.1002/rcm.9358.
- 95. Yang et al., SARS-CoV-2 infection causes dopaminergic neuron senescence, Cell Stem Cell, doi:10.1016/j.stem.2023.12.012.
- 96. **Zavascki** et al., Advanced ventilatory support and mortality in hospitalized patients with COVID-19 caused by Gamma (P.1) variant of concern compared to other lineages: cohort study at a reference center in Brazil, Research Square, doi:10.21203/rs.3.rs-910467/v1.
- 97. **Zeraatkar** et al., Consistency of covid-19 trial preprints with published reports and impact for decision making: retrospective review, BMJ Medicine, doi:10.1136/bmjmed-2022-0003091.
- 98. **Zhang** et al., What's the relative risk? A method of correcting the odds ratio in cohort studies of common outcomes, JAMA, 80:19, 1690, doi:10.1001/jama.280.19.1690.